The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation: Empirical Evidence from Malaysia and Thailand

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Abstract

The purpose of this study is to investigate the effects of oil price fluctuations on exchange rate variations, particularly for Malaysia and Thailand. The study utilizes monthly observation data from 1994 to 2021 and uses structural breaks, cointegration relationships, and Nonlinear Autoregressive Distributed Lag (NARDL) estimates. Both countries had distinct long- and short-term relationships, but there was also an asymmetric relationship between oil prices and multilateral exchange rates revealed by the study. The asymmetric and non-asymmetric causality relationship for both countries indicate a unidirectional non-asymmetric causality between the oil price and the exchange rate variation. The empirical findings will be of great assistance Received: 23.08.2023 Available online: 31.12.2024

to policymakers in evaluating global oil price fluctuations as a crucial indicator for monitoring and stabilizing currencies as a long-term monetary policy strategy.

Keywords: asymmetric, exchange rate, NARD estimates, oil price

JEL: C50, E60

1. Introduction

Crude oil is an important energy source in the world, especially in oil producing countries, as the main source of national income and a catalyst for economic growth and development. The unstable variation of the crude oil price can affect society, producers, manufacturing industries, and the economic performance of a country due to changes in the average spending cost (Cavalcanti and Jalles, 2013). In general, crude oil prices are determined by market forces, and the exchange rate plays a crucial role in determining the value of crude oil, whether it is overvalued, undervalued, or vice versa.

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According to Golub (1983), oil prices influence the real exchange rate of those countries by influencing the asymmetries effect between economic performance and the direction of wealth relocating of the country.

Due to this dynamic relationship, several researchers have attempted to analyze the correlation between oil price changes and exchange rate fluctuations since the 1970 oil price shocks (see Haug and Basher, 2019; Bal and Rath, 2015; Nikbakht, 2010; Chen and Chen, 2007). There are few studies that have demonstrated an existing relationship in terms of cointegration relationship, asymmetric effect, and causality direction (Kisswani, 2021; Basher et al., 2016; Fratzscher et al., 2014; Lizardo and Mollick, 2010; Akram, 2004; Amano and Van Norden, 1998). Malaysia and Thailand were identified by O'Neil et al. (2005) and Burkett and Hart-Landsberg (1998) as the countries that recovered fastest from the Asian Financial Crises in 1997. Moreover, both countries are oil-producing countries in the region and have recorded the same decline trend in export value, production quantity, and local fuel prices. Figures 1 and 2 indicate that the fuel prices of these two countries are markedly different. Figures 3 and 4 show that both countries have distinct oil production capacities from 2010 to 2020. It is undoubtedly true that the oil price shocks have damaged the wave of development and undermined the planned strategies for both countries' economic growth performance, as explained by Kisswani, (2021), Chang, and Wong (2003), and Glasure and Lee (2002).



Figure 1. RON 95 fuel prices trends in Malaysia from 2017 to 2022.





According to Kocaarslan and Soytas (2019), the nonlinearity issue with oil prices may be arising due to data generating issues, regime shifts, and the effect of time-varying coefficients. While Alqaralleh (2020), Brahmasrene et al. (2014), Tiwari et al. (2013), Benhmad (2012), and Akram (2004) postulate that, the nonlinearity of oil price and exchange rate are more likely to stem from the

asymmetric properties of their determinants. By adopting asymmetric responses, both positive and negative aspects in the series can be considered and analysed more effectively. Previous studies have hypothesized that the exchange rate has symmetrically reacted to oil price shocks to gain insight into the impact of oil price variations on economic stability. Therefore, an extensive study should consider





Figure 4. Average oil production of thousands of barrels/per days in Thailand from2010-2020

the asymmetric cointegration and symmetric responses due to the variation of oil prices, which impact exchange rate variation.

The structure of the paper is as follows. Section 2 overviews the literature review of past findings on the relationship between oil prices and exchange rate variations. The data and empirical strategies employed in this study are described in Section 3. Section 4 contains the empirical estimation results and discussions, and Section 5 is where the concluding remarks are presented.

2. Literature Review

The imbalance concentration of surpluses and deficits among the oil exporting countries was significantly influenced by the fluctuation of oil prices, as discussed by Narayan (2013); Lizardo and Mollick (2010). Kisswani et al. (2019) assert that most previous studies that focused on linear analysis have yielded mixed and inconclusive results on the relationship between oil prices and exchange rates. Moreover, Baba and Lee (2022), Bodenstein and Guerrieri (2011), and Kilian (2009) indicate that the fluctuation effect of global oil prices has been possibly driven by different shocks through the international market forces based on demand and supply and trade barrier condition which have resulted from a different impact on the countries macroeconomic variables. The impact of oil price fluctuation on the exchange rate is both positive and negative, depending on the country characteristics and market conditions of Asian countries, as revealed by Nusair and Olson (2019).

During the bullish market for Asian domestic currencies, positive oil prices resulted in an appreciation of the currencies of Indonesia, South Korea, the Philippines, and Thailand. During the bearish market for domestic currencies, positive oil price volatility has resulted in depreciation, which has a constrained effect on Indonesia's currency and an inverse effect on Malaysia's currency. Thus, the currencies responded differently to the oil price fluctuation and had an asymmetrical effect on the bilateral relationship. The relationship between

competitiveness exchange rate and oil prices for oil-exporting countries is influenced by the different exchange rate regime, as shown by Lv et al. (2018). Akram (2004) proved that the relationship between the exchange rate and oil prices in Norwegian samples is influenced by monetary policy, specifically the exchange rate regime.

Moreover, Habid et al. (2016) assert that there is no significant evidence indicating that the oil exporter currencies grew in value compared to the oil importer currencies in the event of a shock in oil prices. Iwayemi and Fowowe (2011), and Rautava (2004) also indicate that there is no significant correlation between oil prices and exchange rates. Huang and Guo (2007) demonstrated that genuine oil price disturbances are only capable of increasing China's long-term competitiveness by a small percentage. Reboredo (2012) has discovered that the relationship between oil prices and exchange rates is usually weak, and it can be a complicated causal relationship that is asymmetric according to the frequency and time frame. A complex causal relationship existed between real oil prices in US dollars, as found by Benhmad (2012), and it shifted over time.

Tiwari et al. (2013) found a causal relationship only at the higher time scale when studying the relationship between oil prices and exchange rates in India. The time-varying causality of these variables, particularly during the crisis, was found by Kocoglu et al. (2022) and Albulescu, and Ajmi (2021), as well. In theory, there are only a few major pathways that could potentially transfer the oil price shock to the exchange rate (Buetzer et al., 2012). First, the trade-in goods are the primary focus of the term of trade, while non-traded goods are not. Additionally, changes to portfolio allocations are necessary to rectify

The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation

the effects of wealth, trade balances, and portfolio allocation related to wealth.

Several studies concluded a theoretical transmission mechanism between oil prices and exchange based on different channels; the term of the trade (Chen and Chen, 2007; Amano and Van Norden, 1998; Corden and Neary 1982), the term of wealth effects (Golub, 1983) and asset pricing theory (Chen et al., 2010; Akram, 2009; Frankel, 2006). Zhang et al. (2016) determined through the transmission mechanism that active financial activities determine oil prices and the exchange rate, and these changes can be brief and swift in the market. Similarly, Beckmann et al. (2020) have concluded that the shortterm bidirectional relationship between those variables is bidirectional. Baghestani and Toledo (2019) assume that the changes in oil prices accurately predict the direction of the exchange rate more than 3 months ahead when studying NAFTA countries. Lizardo and Mollick (2010) also found that real oil prices are the most important factor in predicting future exchange rate movements. Habid and Kalamova (2007) and Cashin et al. (2004) discovered that few countries studied have a significant connection between oil prices and competitiveness of exchange rates. Furthermore, Chen et al. (2010) validated the argument by demonstrating that commodity prices are competitive on the global market and have varied impacts on all countries, regardless of whether they are oil exporters or importers.

Gao et al. (2022) have revealed a positive correlation between oil prices and exchange rates, with a mixed causality effect for South Asian countries. In addition, Kisswani et al. (2019) also demonstrated a mixed causality, and Malaysia has experienced an asymmetrical impact. According to Alqaralleh

Author(s)	Time Span	Country	Estimation technique	Causality direction
Wang et al. (2022)	2008-2021	Belt and Road countries	EMD	OP cause ER
Butt et al. (2020)	1994-2017	Malaysia	TVECM	OP cause ER
Haug and Basher (2019)	1976-2014	Oil exporting countries ARCH-GARCH		No causality
Singhal et al. (2019)	2006-2018	Mexico	ARDL	OP cause ER
Tran et al. (2019)	2005-2017	ASEAN countries	Markov Switching model	OP cause ER
Bhattacharya et al. (2019)	2000-2016	India	DCC-GARCH	OP cause ER
Ji et al. (2019)	2011-2017	China and US	CO-VAR	OP cause ER
Hussain et al. (2017)	2006-2016	Asian countries	DCCA	Weak relationship
De Vita and Trachanas (2016)	1994-2013	China and India	NARDL	Weak relationship
Bal and Rath (2015)	1994-2013	China and India	VAR-GARCH	Depend on country situation
Nikbakht (2010)	2000-2007	OPEC countries	Panel estimates	ER cause OP
Chen and Chen (2007)	1972-2005	G7 countries	Panel estimates	ER cause OP

Table 1.	Selected	studies or	n causality	relationship	between o	il prices and	exchange rate
							J

Note: ER refer to exchange rate and OP indicate the oil price.

(2020), G-20 countries experience a mixed reaction to exchange rates and oil price fluctuations. The short-period fluctuation effect of oil prices has been validated by Gao et al. (2017), which shows a diverse dynamic characteristic of exchange rates for short-run periods. Thus, the risk associated with oil price fluctuation becomes greater, creating a complicated situation in forecasting future oil price fluctuations and relationships with other indicators. Zhang et al. (2018) provided evidence of the country's unique features and the time-series pattern that created a fluctuation in natural energy prices in Japan, the United States, and Germany. In summary, most previous studies have examined the relationship between exchange rates and oil prices, whether they are bilateral or multilateral exchange rates, by utilizing various econometric estimations. Although the primary finding was that exchange rates and oil prices had a relationship, the relationship results varied depending on the time frame, countries' characteristics, and other external factors. The following Table 1 shows the causal relationship between oil prices and exchange rates from various studies worldwide:

3. Methodology

This study uses monthly data covering 1994 until 2021 based on the available data for both Malaysia and Thailand. The data are composed of the multilateral exchange rate and the global price of Brent crude oil, which

is measured as OPEC (\$/bbl.) and has a constant price in 2000. The study determines the multilateral exchange rate by combining the weighted average of the real exchange rates of all trading partners in the country and comparing it to the indication of the real effective exchange rate (REER) in the domestic currency. The appreciation of the domestic currency and its purchasing power and competitiveness against trading partners' currencies can be observed through an increased value in the multilateral exchange rate. Data on multilateral exchange rates is available in the JP Morgan Database (2022). While the Bank of Thailand (2022), Ministry of Finance Malaysia (2022), and the International Monetary Fund (2022) provide data on oil prices.

To ensure that the variables used are free of stationary problems, we impose the unit root test as a first step. We carry out the Augmented Dickey-Fuller test (ADF) (1981) and the endogenous structural breaks based on Zivot and Andrew (ZA) (2002) unit root test since there is a multilateral exchange rate effect for Malaysia and Thailand which might have structure breaks consequence of changes in the exchange rate regime from the pegged regime to managed floating regime. The nonlinear stationarity effect with the tested variables in this study is captured using the nonlinear Fourier ADF unit root test.

To measure the long-term relationship between oil prices and the multilateral exchange rate during the sub-period of this study, it is necessary to apply the Gregory and Hansen (1996) cointegration test. The null hypothesis of no cointegration is rejected by this approach when the extension test of the estimation types is dependent on the existence of a cointegration relationship with the subsample. The GH cointegration test suggested The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation

three model specification alternatives that were able to withstand structural breaks that rejected the business's probability.

First, the level shift model only allows changes in the intercept, as shown in equation (1).

$$y_{1t} = \mu_1 + \mu_2 \varphi_{t\tau} + \alpha' y_{2t} + \varepsilon_t \qquad t = 1, \dots, n$$
(1)

Secondly, the trend model accompanies the trends in the data and at the same time prevents a change in the level as presented in the equation (2).

$$y_{1t} = \mu_1 + \mu_2 \varphi_{t\tau} + \beta_t + \alpha' y_{2t} + \varepsilon_t$$
$$t = 1, \dots, n$$
(2)

Thirdly, the regime model specification allows for changes in both the intercept and slope of the cointegration vector, as shown below.

$$y_{1t} = \mu_1 + \mu_2 \varphi_{t\tau} + \alpha' y_{1t} + \alpha'_2 y_{2t} \varphi_{t\tau} + \varepsilon_t$$
$$t = 1, \dots, n$$
(3)

Additionally, the dummy variable mentioned in the equations, which captures the structural alteration, can be presented as follows:

$$\varphi_{t\tau} = \begin{cases} 0, & t \le [n\tau] \\ 1, & t > [n\tau] \end{cases}$$
(4)

The $\varphi_{t\tau}$ and y_{2t} , are variables that are stationary at l(1), ε_t is an error term at the l(0), μ_n and α' are time-invariant and $\tau \varepsilon(0, 1)$ is the changes point of relative timing, which is normally taken to be (0.15n, 0.85n) interval. This GH cointegration approach computed the usual ADF and PP test statistics, denotes as $ADF(\tau)$, $Z_{\alpha}(\tau)$ and $Z_{t}(\tau)$ for the possible existing breakpoints and choosing the smallest value which represents greater evidence in the long-run cointegration condition. Since the variables used in this study tend to have

nonlinearities in the long-run relationship, we therefore implement the NARDL estimation proposed by Shin et al. (2014) using the following equations:

$$iv_{t}^{+} = \sum_{j=1}^{t} \Delta iv_{j}^{+} = \sum_{j=1}^{t} max \ (\Delta iv_{j}, 0) \quad (5)$$
$$iv_{t}^{-} = \sum_{j=1}^{t} \Delta iv_{j}^{-} = \sum_{j=1}^{t} min \ (\Delta iv_{j}, 0) \quad (6)$$

A general form for NARDL estimation specification can be written as:

$$\Delta REER_{t} = \alpha_{0} + \beta_{0}REER_{t-1} + \beta_{1}Oil_{t}^{+} + \beta_{2}Oil_{t}^{-} + \sum_{p} \mu_{i} \Delta REER_{t-1} + \sum_{i=1}^{Q} (\gamma_{i}^{+}OIL_{t-i}^{+} + \gamma_{i}^{+}Oil_{t-i}^{-}) + \varepsilon_{t}$$
(7)

where, REER is the dependent variable, Oil is the independent variable and the parameters p and q are the lag order. The NARDL estimation approach allows a simultaneous cointegration model for a long and short-run relationship with asymmetric relationship which accounted for the different decomposition effects of positive and negative shocks of oil prices on multilateral exchange rates and vice versa. In addition, the NARDL approach can accommodate any order of integration, whether it's I(0), I(1), or a combination of both.

We underscored the significance of asymmetric non-causality in determining the causal direction of oil price shocks, which impact the exchange rate and vice versa (see Nouira et al., 2019; and Hatemi-J et al., 2016). To account for Hatemi-J non-causality, the cumulative sums of positive and negative components of the underlying variables were used, using bootstrap critical values. Hatemi-J variable, v is the (1x2) vectors of intercepts,

(2012) suggested the equation that forms the asymmetric non-causality estimates.

$$y1_t = y1_{t-1} + \varepsilon 1_t = y_{10} + \sum_{i=1}^{t} \varepsilon 1_i \qquad t = 1,2 \dots T$$
(8)

$$y_{2_t} = y_{2_{t-1}} + \varepsilon_{2_t} = y_{20} + \sum_{i=1}^t \varepsilon_{2_i} \qquad t = 1, 2 \dots T$$
(9)

where, the y_{10} and y_{20} are the constant's values; and the $\varepsilon 1_i$ and $\varepsilon 2_i$ are white noise disturbance term, $\varepsilon_{1i}^+ = \max(\varepsilon_{1i}, 0),$ $\varepsilon_{2i}^{+} = \max (\varepsilon_{2i}^{-}, 0), \quad \varepsilon_{1i}^{-} = \min (\varepsilon_{1i}^{-}, 0),$ and $\varepsilon_{2i}^{-} = \min (\varepsilon_{2i}^{-}, 0)$ represent the positive and negative parameters (also can be expressed as $\varepsilon \mathbf{1}_i = \varepsilon_{\mathbf{1}i}^+ + \varepsilon_{\mathbf{1}i}^ \varepsilon 2_i = \varepsilon_{2i}^+ + \varepsilon_{2i}^-$. This condition can be described as follows:

$$y1_{t} = y1_{t-1} + \varepsilon 1_{t} = y_{1,0} + \sum_{i=1}^{t} \varepsilon_{1i}^{+} + \sum_{i=1}^{t} \varepsilon_{1i}^{-} (10)$$
$$y1_{t} = y2_{t-1} + \varepsilon 2_{t} = y_{2,0} + \sum_{i=1}^{t} \varepsilon_{2i}^{+} + \sum_{i=1}^{t} \varepsilon_{2i}^{-} (11)$$

where, the positive and negative values for each variable be defined in a cumulative form as $y_{1t}^+ = \sum_{i=1}^t \varepsilon_{1i}^+$, $y_{1t}^- = \sum_{i=1}^t \varepsilon_{1i}^-$ and $y_{2t}^+ = \sum_{i=1}^t \varepsilon_{2i}^+$, $y_{2t}^- = \sum_{i=1}^t \varepsilon_{2i}^-$. Hence, the underlying variables have a permanent impact on positive and negative shocks, respectively. The causal relationship between variables is tested by using the VAR model with order p-value:

$$y_1^+ = v + A_1 y_{t-1}^+ + \dots + A_p y_{t-1}^+ + \mu_t^+;$$

where, $y_t^+ = (y_{1t}^+, y_{2t}^+)$ (12)

$$y_{1}^{-} = v + A_{1}y_{t-1}^{-} + \dots + A_{p}y_{t-1}^{-} + \mu_{t}^{-};$$

where, $y_{t}^{+} = (y_{1t}^{-}, y_{2t}^{-})$ (13)

where, then $y_{t}^{+/-}$ is the (1x2) vectors of

 $\mu_t^{+/-}$ is the (1x2) vectors of error terms and A_r is the (2x2) matrix of parameters for lag order *r*=1,, *p*, and the lag order selection is based on the Hatemi-J criterion (HJC).

4. Empirical Results

To determine their level of stationarity, all variables in the study were subjected to multiple unit root tests. Table 2 presents the unit root test statistics for all variables. The results of the traditional and endogenous structural breaks unit root tests proved the unit root hypothesis as I(1) for all variables, as well as the existence of a structural break in the series of variables. Between 1998 and 2008, the structural break showed the effect of variation during the Asian Financial The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation

Crisis and the Global Financial Crisis. Most economies in the world have been affected by both financial crises, particularly in emerging markets in East and Southeast Asia countries during the late 1990s and 2000s. The Fourier ADF unit root test reveals that the REER variables in Thailand are experiencing a unit root issue. Utilizing different analysis approaches is necessary due to the integration of unit root test statistics as I(1) and mixed, as demonstrated in Table 3.

Table 4 shows the estimates of long-run cointegration relationships for the sub-period of a series based on the GH estimates. The statistical result for Malaysia shows a rejected null hypothesis of no cointegration for the regime and regime trend estimation model,

	At level		At first difference		
	ADF	ZA	ADF	ZA	
REER _{Mal}	-2.299	-3.197 (2015/M06)	-17.177***	-17.610*** (1998/M02)	
Oil _{Mal}	-1.718	-4.423 (2014/M08)	-17.063***	-17.259*** (2008/M08)	
REER _{Thai}	-2.299	-3.066 (1998/M11)	-17.177***	-17.344*** (1997/M12)	
Oil _{Thai}	-2.021	-2.801 (2002/M02)	-13.917***	-8.872*** (1999/M01)	

Table	2.	Unit	root	test	results
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Note: *** denotes rejection of the null of a unit root at the 1% level.

 Table 3. Fourier ADF unit root test results

	Min. SSR	Ř	Lag	FADF-stat	F()
REER _{Mal}	0.023	2	1	-2.443	1.068
Oil _{Mal}	0.609	1	3	-2.797	3.082
REER _{Thai}	0.032	1	5	-4.721***	5.778***
Oil _{Thai}	0.609	1	1	-2.797	3.082

Note: The finite sample critical values for FADF and F-test are taken from Table 1 of Enders and Lee (2012). *** Denotes rejection of the null of a unit root at the 1% level.

Estimations	ADF-stat	Т _в	Z _t -stat	T _B	Z₂-stat	T _B		
Malaysia								
Level shift	-4.20	1998/M06	-4.31	1998/M01	-31.21	1998/M01		
Level shift with trend	-3.59	1999/M10	-4.75*	1999/M06	-34.36	1999/M06		
Regime change	-6.09***	1998/M03	-5.94**	1997/M07	-61.69**	1997/M01		
Thailand								
Level shift	-5.09**	2002/M06	-4.39*	2002/M03	-30.01	2002/M03		
Level shift with trend	-4.29	2002/M06	-4.53	2003/M11	-31.77	2003/M11		
Regime change	-5.99**	2004/M02	-5.90**	2003/M11	-49.67	2003/M11		

Table 4. Gregory and Hansen cointegration test results

Note: ***, ** and *refer to significant value of 1, 5 and 10% based on Gregory and Hansen (1996) critical value statistics. TB represent the break date from the GH cointegration estimation.

with break dates occurring between 1997 and the end of 1999. Thailand's results rejected the null hypothesis that a level and regime trend with a break date in 2002 and 2004 did not have cointegration. The break dates in line with the Asian Financial Crises have worsened both countries' economic progress, highlighted by Nambiar (2003) and Chin and Jomo (2001). Political issues and changing monetary policy caused the loosening of credit extensions and the reduction of interest rates in 2002, leading to the expectation of break-up dates in Thailand (Mutebi, 2003). With the Royal Thai Army coup against the ruling government, Thailand's political issues become more heated, resulting in a shortterm currency depreciation shock in 2006 (Reuters, 2008).

According to a study by Bahmani-Oskooee and Baek (2017), a nonlinear model was superior to a linear model in identifying an exchange rate. The asymmetric effect of oil prices on the multilateral exchange rate was investigated by this study using the NARDL approach, and the statistical results are presented in Table 5. By setting out lag orders, this study used the general-specific approach to select actual NARDL estimations.

The reason for eliminating insignificant regressors is because they can cause estimation inaccuracies and generate noise in the dynamic multiplier. The critical upper bound has been exceeded in Malaysia, which suggests that oil prices and the multilateral exchange rate have an asymmetric long-run relationship. Thailand, despite being at the 10% significant critical value between bound critical values, recorded an inconclusive result, which is surprising. While the $t_{_{\rm RDM}}$ statistics value proposed by Baneriee et al. (1998) in detecting the asymmetric cointegration is significant in Malaysia and not indicating a significant result for Thailand; and this finding was almost similar to Kisswani et al. (2019) who presented a similar result while studying the impact of oil price on the exchange rate for ASEAN-5.

This study continues to test dynamic asymmetric relationships in both the long and short-run terms as shown in Table 5 to confirm the appropriateness of performing the asymmetric model. There is an existence of asymmetric relationships across both positive and negative sides, and the oil price has recorded a significant level of relationship for both positive and negative

long-run relationships with the estimated coefficients at 0.073 and -0.100 respectively for Malaysia's multilateral exchange rate. The multilateral exchange rate indicator shows an almost 7.3% increase in response to a positive 1% increase in world oil prices, as can be assumed analytically. Instead, a negative 1% decrease in the world oil price level leads to an almost 10% depreciation in the multilateral

The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation

exchange rate indicator. Furthermore, the long-term relationship between oil prices and Malaysia's multilateral exchange rate indicator is asymmetric. The oil price fluctuation in Thailand resulted in an insignificant coefficient value for both positive and negative long-run estimates. In the short run, the oil prices variable has an asymmetrical relationship with Thailand's multilateral exchange rate.

	Mala	aysia	Thailand		
	Coef.	Std. error	Coef.	Std. error	
REER ⁺	-0.078*** (0.000)	0.019	-0.048*** (0.004)	0.016	
Oil ⁺ t	0.006* (0.052)	0.003			
Oil _t	0.008** (0.027)	0.004			
$\Delta \text{REER}^{+}_{t-1}$			0.119** (0.039)	0.057	
$\Delta \text{REER}^+_{t=3}$	0.071** (0.040)	0.058	-0.125** (0.030)	0.057	
$\Delta \text{REER}^+_{t-4}$			0.145** (0.013)	0.058	
$\Delta \text{REER}^+_{t-7}$			-0.111* (0.056)	0.057	
$\Delta 0il_{t-3}^+$			-0.078** (0.011)	0.030	
∆0ilī	0.054*** (0.006)	0.019			
$\Delta 0il_{t-1}^{-}$	-0.034* (0.060)	0.018			
$\Delta 0il_{t-2}^{-}$			-0.039* (0.076)	0.022	
$\Delta 0il_{t-3}^{-}$	-0.032* (0.056)				
LR ⁺	0.073** (0.037)		-0.024 (0.734)	0.116	

Table 5. The NARDL estimation results

	Mala	iysia	Thailand				
	Coef.	Std. error	Coef.	Std. error			
LR⁻	-0.100** (0.014)		0.045 (0.591)	0.289			
LR-asymmetric	16.14*** 2.017 (0.000) (0.157)						
SR-asymmetric	2.5 (0.1	39 12)	4.967** (0.027)				
F _{PSS}	5.396**						
t _{BDM}	-3.989**						
Diagnostic Test							
Serial Correlation	3.690	0.145	5.990	0.245			
Breusch-Pagan	0.330	0.330 0.567 0.321		0.487			
D-W Statistics	1.9	1.903 1.807					

Note: Only the significant coefficient reported through the estimation results and values in () indicate the p-value. ***, ** and * denote rejection of the null hypothesis of no cointegration at the 1, 5 and 10% significance levels.

The long-term asymmetric relation with oil price fluctuation revealed by Butt et al. (2020) and Kisswani et al. (2019) in Malaysia is akin to our findings. The study by Bahmani-Oskooee and Kanitpong (2019) and Ibrahim and Chancharoenchai (2013) has demonstrated a short-run asymmetric relationship between oil prices and economic growth in Thailand. Additionally, both Sanusi (2020) and Basnet and Upadhyaya (2015) came to the conclusion that there is no clear relationship between the global oil price fluctuation and exchange rate for ASEAN countries, as well as oil exporting and oil-importing countries. Basnet and Upadhyaya (2015) argue that oil prices do not induce fluctuations in the exchange rate of ASEAN countries, and some ASEAN countries

exhibit a unique pattern of response to global oil price fluctuations. The exchange rate's impact on oil prices is different between oilproducing and non-oil-producing countries, as revealed by Wang et al. (2022).

Figure 5 demonstrates the long-term dynamic multiplier effects of positive and negative changes in world oil prices in Malaysia and Thailand. The oil price in Malaysia experienced a fluctuating long-run cumulative pattern in the early period, then dropped and then crossed over to negative changes approximately in the 10-period and remained until the end period. Thailand's outcome fluctuates during the early estimation period, then rises and overshoots the positive, and remains until the end of the study period.

The Asymmetric Relationship Between Oil Price Fluctuation on Exchange Rate Variation

Articles



Cumulative effect of LNOIL on LNRER





Cumulative effect of LNOIL on LNRER

Figure 5(b). The cumulative multiplier effects for Thailand

Table 6. Asymmetric and non-asymmetric causality results.

Null hypothesis	Mala	ysia	Thailand					
Null hypothesis	Test value	Causality	Test value	Causality				
Asymmetric causality								
$OiI - / \rightarrow REER$	0.113	No	0.340	No				
Oil ← / – REER	0.007	No	0.235	No				
Non-asymmetric causality	Non-asymmetric causality							
$Oil^+ - / \rightarrow REER^+$	3.991***	Yes	4.001***	Yes				
$Oil^+ \leftarrow / - REER^+$	0.211	No	0.131	No				
$OiI^ / \rightarrow REER^-$	0.433	No	0.007	No				
Oil ⁻ ← / – REER ⁻	1.123	No	0.032	No				

Note: *** denote rejection of the null hypothesis of causality at the 1% significance level based on the bootstrap critical value.

We tested the possibility of an asymmetric causality relationship between oil prices and multilateral exchange rates using the Hatemi-J (2012) nonlinear approach. The bootstrap critical value was used to determine positive and negative changes in asymmetric non-causality using the Hatemi-J (2012) Gauss code, as shown in Table 6. Based on the results presented, it is unclear whether changes in oil prices, whether rising or falling in the symmetric condition, have any causality direction for both countries. Both countries experience strong positive causal effects from oil price fluctuations to exchange rate variations. This outcome parallels those of Khraief et al. (2021) and Jung et al. (2020) because of the trade balances between these two countries, which are among the biggest oil producing nations in the ASEAN region.

The non-asymmetric causality between oil prices and REER in Malaysia was caused by the fluctuating value of the US dollar, which is the invoice currency for international trade, in comparison to the use of local currency. The value of the local currency will be depreciated and the value of the local

multilateral exchange rate will be directly reduced by an appreciation of the US dollar, and vice versa. Wang et al. (2022) in a recent study found that the relationship for long cycle periods was influenced by the exchange rate regime, balance of payment, and economic performance. The oil price shock has had a negative impact on China's exchange rate, as evinced by Ju et al. (2014), despite having a positive impact on China's GDP.

5. Concluding Remarks

The results of this study demonstrated an asymmetric and causal connection between exchange rate variation and oil price fluctuation, especially in light of the dataset used in the study. The non-asymmetric causality exhibits a one-way causality between oil price fluctuation and exchange rate variation, giving a clear picture of both countries' effects on unstable global oil prices. The NARDL estimates revealed that the impact of oil price appreciation on exchange rate variability was distinct from the impact of oil price depreciation on multilateral exchange rates that involved both countries. Policymakers

may need to rethink the monetary and fiscal framework to reduce the financial burden due to the uncertainties of global commodity prices, inflationary pressure, currency volatility and instability, and economic growth. The exchange rate for both countries is primarily influenced by the fluctuation of oil prices. By considering the expansion of petrol and export subsidy programs, particularly for local retail consumption, policymakers can mitigate the risk of shocks. Financial instruments that are suitable for absorbing exchange rate transactions and translation risk are required. Future monetary policies can be improved through monitoring currency movements and the social well-being of society in light of the global oil price shock, as indicated by the asymmetric effects.

Acknowledgement

We want to express our gratitude to UTM for granting the UTM Fundamental Research Fund (Vote Number: Q.J130000.3855.21H93).

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